# §15. Exponential Random Variables

Definition 1 A exponential random variable  $X$  has pdf:  $p(x) = \begin{cases}$  $\lambda e^{-\lambda(x-a)}$  ;  $x \ge a$  ,  $\lambda > 0$ 0 ; otherwise  $X \sim \exp(\lambda, a)$ 



Figure 1: Graph of Exponential Distribution

Let's compute the expected value, variance, and cumulative distribution.

$$
E(X) = \int_{a}^{\infty} x \cdot \lambda e^{-\lambda(x-a)} dx = x \cdot \lambda \cdot \frac{e^{-\lambda(x-a)}}{-\lambda} \Big|_{a}^{\infty} - \int_{a}^{\infty} \lambda \cdot \frac{e^{-\lambda(x-a)}}{-\lambda} dx
$$

$$
= a + \frac{e^{-\lambda(x-a)}}{-\lambda} \Big|_{a}^{\infty}
$$

$$
= a + \frac{1}{\lambda}
$$

$$
\operatorname{Var}(X) = \int_a^\infty x^2 \cdot \lambda e^{-\lambda(x-a)} dx - \left(a + \frac{1}{\lambda}\right)^2 = \left(a^2 + \frac{2a}{\lambda} + \frac{2}{\lambda^2}\right) - \left(a + \frac{1}{\lambda}\right)^2
$$

$$
= \frac{1}{\lambda^2}
$$

$$
\sigma(X) = \frac{1}{\lambda}
$$

$$
F(x) = \int_{a}^{x} \lambda e^{-\lambda(x-a)} dt = \lambda \cdot \frac{e^{-\lambda(x-a)}}{-\lambda} \bigg|_{a}^{x} = 1 - e^{-\lambda(x-a)} \quad , \quad x \ge a
$$



Figure 2: Graph of Cumulative Distribution

#### Example 1

The waiting time for the secretary of a family physician to pick up the phone from the automatic system is exponential  $X \sim \exp\left(\frac{1}{4}\right)$  $(\frac{1}{4}, 0)$ where time is in minutes.

- a. What is the expected waiting time before the secretary picks the phone and what is the standard deviation?
- b. What is be probability that the patient has to wait mare than 10 minutes ?

Solution

a. 
$$
E(X) = a + \frac{1}{\lambda} = 0 + \frac{1}{1/4} = 4 \text{ min}, \qquad \sigma(x) = \frac{1}{\lambda} = \frac{1}{1/4} = 4 \text{ min}
$$
  
b.  $P(X > 10) = 1 - F(10) = 1 - (1 - e^{-\frac{1}{4}(10 - 0)}) = e^{-2.5} = 0.0821$ 

Note that: If  $X \sim \exp(\lambda, a)$  then

$$
P(X \ge s + t | X \ge s) = \frac{1 - (1 - e^{s + t - a})}{1 - (1 - e^{-\lambda(s - a)})} = \frac{e^{-\lambda(s + t - a)}}{e^{-\lambda(s - a)}} = e^{-\lambda t}
$$

$$
P(X \ge t + a) = 1 - (1 - e^{-\lambda(t + a - a)}) = e^{-\lambda t}
$$

Remember, a is the initial time, so the probability that a patient will have to wait t minutes more given s minutes have already passed waiting is the same as waiting for  $t$  minutes initially. This is the **memoryless** property of the exponential distribution.

### Relation with a Poisson process

Let Y be a Possion process with rate  $\lambda$ . Then in time t we expect  $\lambda t$ events. What is the probability that we have to wait at least  $x$  until the  $1<sup>st</sup>$  Poisson event. Let X be the random variable which is the waiting time until  $1^{st}$  Poisson event.

$$
P(X > x) = P(Y = 0) = \frac{e^{-\lambda x}(\lambda x)^0}{0!} = e^{-\lambda x}
$$

$$
F(x) = P(X \le x) = 1 - e^{-\lambda x}, \qquad x \ge 0
$$

Example 2

A radioactive source is emitting  $\alpha$ -particles as a Poisson process at a rate of 8 particles per minute. What is the probability that the interval between the emissions of two particles will be more than 40 sec?

$$
P\left(X > \frac{2}{3}\right) = 1 - F\left(\frac{2}{3}\right) = 1 - \left(1 - e^{-8 \cdot \frac{2}{3}}\right) = 0.00483
$$

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## Example 3

Let  $X$  denote the time between detections of a particle with a Geiger counter and let x be exponential with  $E(x) = 0.125$  min

- a. What is the probability that we will detect a particle within the first 10 sec ?
- b. A minute has passed without detecting a particle. What is he probability that we will detect ore within the next 10 sec?

## Solution

a. 
$$
P(x < \frac{1}{6}) = F(\frac{1}{6}) = 1 - e^{-\frac{1}{6 \cdot 12} 5 \cdot \frac{1}{6}} = 1 - e^{-\frac{8}{6}} = 0.7364
$$

b.

$$
P\left(x < 1\frac{1}{6} \mid x > 1\right) = \frac{p\left(1 < x < 1\frac{1}{6}\right)}{P(x > 1)} = \frac{F(1\frac{1}{6}) - F(1)}{1 - F(1)}
$$
\n
$$
= \frac{\left(1 - e^{-8.7/6}\right) - \left(1 - e^{-8}\right)}{1 - \left(1 - e^{-8}\right)}
$$
\n
$$
= \frac{e^{-8} - e^{-8.7/6}}{e^{-8}}
$$
\n
$$
= 1 - e^{8.1/6}
$$
\n
$$
= 0.7364
$$

