# §16. Gaussian Random Variables

**Definition 1** 

A random variable X is **Gaussian (normal)** if it has the pdf

$$P(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \qquad -\infty < x < \infty, \qquad \sigma > 0$$

and  $X \sim N(\mu, \sigma)$ 

Let's check that this is a valid pdf.

Let 
$$z = \frac{x-\mu}{\sigma}$$
. Then  
$$\int_{-\infty}^{\infty} p(x) dx = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz = I$$

Let's rotate  $\frac{1}{\sqrt{2\pi}}e^{-\frac{z^2}{2}} = w$  about the *w*-axis.



The resulting surface is  $w = \frac{1}{\sqrt{2\pi}}e^{-\frac{z^2+y^2}{2}}$ 

The volume under this surface is:

$$\text{Vol.} = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-\frac{z^2 + y^2}{2}} dz \, dy = \frac{1}{\sqrt{2\pi}} \left( \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx \right) \left( \int_{-\infty}^{\infty} e^{-\frac{y^2}{2}} dy \right) = I^2 \sqrt{2\pi}$$



Figure 1: Normal Distribution. Observe the very thin tails

Using cylindrical shells

$$\begin{aligned} \text{Vol.} &= \frac{1}{\sqrt{2\pi}} \int_0^\infty 2\pi r e^{-\frac{r^2}{2}} dr = \sqrt{2\pi} \int_0^\infty r e^{-\frac{r^2}{2}} dr = \sqrt{2\pi} \\ \sqrt{2\pi} &= I^2 \sqrt{2\pi} \quad \Rightarrow \quad I^2 = 1 \quad \Rightarrow \quad I = 1 \quad ; \quad \text{this is a valid density.} \end{aligned}$$

Next, let's compute the mean and variance

$$\begin{split} E(X) &= \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \qquad \text{Let } z = \frac{x-\mu}{\sigma} \quad , \quad x = \mu + z\sigma \\ &= \mu \cdot \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz + \sigma \cdot \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z e^{-\frac{z^2}{2}} dz \quad 0 \\ &= \mu \end{split}$$

$$\begin{split} E(X^2) &= \int_{-\infty}^{\infty} x^2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \qquad , \quad z = \frac{x-\mu}{\sigma} \\ &= \int_{-\infty}^{\infty} \frac{(\sigma z + \mu)^2}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz \\ &= \underbrace{\sigma^2 \int_{-\infty}^{\infty} \frac{z^2}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz + 2\mu\sigma \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz + \mu^2}_{integration \ by \ parts = 1} \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz}_{1} \\ &= \sigma^2 + \mu^2 \end{split}$$

$$\operatorname{Var}(X) = E(X^2) - E(X)^2 = \sigma^2$$

**Theorem 1**  
If 
$$X \sim N(\mu, \sigma)$$
 then  $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$ 

Proof.

$$F(z) = P(Z \le z) = P\left(\frac{X-\mu}{\sigma} \le Z\right) = P(X \le \mu + z\sigma) = F(\mu + z\sigma)$$
  
Remember that  $F(x) = \int_{-\infty}^{z} p(t) dt \implies F'(z) = p(z)$ 

$$P(z) = F'(z) = F'(\mu + z\sigma) \cdot \sigma = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(\mu + z\sigma - \mu)^2}{2\sigma^2}} \cdot \sigma$$
$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

Thus,  $Z = \frac{X - \mu}{\sigma}$  is standard normal.

The **error function** is defined as

$$\operatorname{erf}(X) = \frac{2}{\sqrt{\pi}} \int_0^x e^{-t^2} dt$$

The **cumulative** of the standard normal is

$$\Phi(Z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{z} e^{-t^2} dt = \frac{1}{2} + \frac{1}{2} \operatorname{erf}\left(\frac{z}{\sqrt{2}}\right)$$

Example 1

- a. P(Z < -1.82) = 0.0344
- b. P(Z < 1.09) = 0.8621
- c.  $P(Z > 0.64) = 1 P(Z \le 0.64)$ = 1 - 0.7389 = 0.2611
- d.  $P(0.12 < Z \le 2.08) = P(Z \le 2.08) P(Z < 0.12) = 0.9812 0.578 = 0.4334$

# Example 2

The student grades in Calculus 2 are normally distributed with  $\mu = 68$  and  $\sigma = 13$ . What is the probability that the Cal 2 grade of a random student is

- a. X > 90
- b. X < 40
- c. 70 < X < 75

Solution

a. 
$$Z = \frac{90 - 68}{13} = 1.69$$
  
 $P(X > 90) = P(Z > 1.69)$   
 $= 1 - 0.9545$   
 $= 0.0455$   
b.  $Z = \frac{40 - 68}{13} = -2.15$   
 $P(X < 40) = P(Z < -2.15)$   
 $= 0.0158$ 



Figure 2: Graph of  $\operatorname{erf}(X)$ 

c. 
$$Z = \frac{70.68}{13} = 0.15$$
  $Z = \frac{75 - 68}{13} = 0.54$ 

$$P(70 < X < 75) = P(0.15 < Z < 0.54)$$
$$= 0.7054 - 0.5596$$
$$= 0.1458$$

## Example: Cont'd

X = 75  $P \stackrel{?}{=} 0$ . We are approximating a discrete distribution with continuous. To be discussed in detail in the next lecture.

## Example: Cont'd

What are the top 10% and bottom 10% percentiles of the Cal 2 grades distribution.

 $P_{10}: \quad Z = -1.28 \qquad x = \mu + z\sigma = 68 + (-1.28)13 = 51.36$ 

 $P_{90}: Z = 1.28$   $x = \mu + z\sigma = 68 + (1.28)13 = 84.64$ 

## Example 3

The average weight of Canadian men is  $\mu = 178.2 \, cm$  with  $\sigma = 7.4 \, cm$ . Assume Gaussian distribution of heights.

- a. What is the probability that a randomly selected Canadian man is more than 2m tall?
- b. J.T. is 188.0 cm tall. At what percentile of the Canadian male population is his height?
- c. What is the top 20% percentile of the height distribution of Canadian males?

Solution

a. 
$$Z = \frac{200 - 178.2}{7.4} = 2.95$$
  
 $P(X > 200) = P(Z > 2.95)$   
 $= 1 - 0.9984$   
 $= 0.6016$ 

b. 
$$Z = \frac{188 - 178.2}{7.4} = 1.32$$
  
 $P(Z < 1.32) = 0.9066 \sim P_{91}$ ; The ninety-first percentile  
c.  $P_{80} \rightarrow Z = 0.84$   
 $x = \mu + z\sigma = 178.2 \pm 0.84(7.4) = 184.42 \, cm.$ 

Note: That for any Gaussian distribution we have:

 $P(\mu - \sigma \le X \le \mu + \sigma) = P(-1 \le z \le 1) = 0.6827$ Chebychev says nothing  $P(\mu - 2\sigma \le X \le \mu + 2\sigma) = P(-2 \le z \le 2) = 0.9545$ At least  $1 - \frac{1}{2^2} = 0.75$  $P(\mu - 3\sigma \le X \le \mu + 3\sigma) = P(-3 \le z \le 3) = 0.9973$ At least  $1 - \frac{1}{3^2} = 0.8889$